

Appendix B – Stochastic partial differential equation approximation to Gaussian random fields

As a maximum likelihood implementation of the spatial Gompertz model, we use a stochastic partial differential equation (SPDE) approximation to a Gaussian random field (Lindgren et al. 2011), as previously implemented and tested in the integrated nested Laplace approximation (INLA) software (Rue et al. 2009). This approach approximates a Gaussian random field using a Gaussian Markov random field whose pairwise correlations follow the Matérn class. This approximation yields the following equivalence:

$$\mathbf{Q} = \tau^2(\kappa^4 \mathbf{C} + 2\kappa^2 \mathbf{G}_1 + \mathbf{G}_2) \quad (\text{B.1})$$

where \mathbf{Q} is the precision matrix of Gaussian Markov random field approximation, κ and τ are parameters in the Matérn approximation (when specifying Matérn smoothness parameter $v=1$), and \mathbf{C} , \mathbf{G}_1 , and \mathbf{G}_2 are sparse matrices representing a piecewise linear basis function for the approximation (see Lindgren et al. 2010, Lindgren and Rue 2013 for details). \mathbf{C} , \mathbf{G}_1 , and \mathbf{G}_2 are calculated using the R-INLA software (Lindgren and Rue 2013) in two steps. First, nodes for a finite element analysis “mesh” are calculated using R-INLA, where this mesh defines a piecewise linear (i.e., triangular in 2-dimensional space) approximation to \mathbf{C} , \mathbf{G}_1 , and \mathbf{G}_2 between nodes. This mesh has K nodes, where nodes are included at each of I stations as well as additional locations, and the number of additional locations can be predefined to control the tradeoff between precision and computational complexity of the SPDE approximation. R-INLA then calculates values for \mathbf{C} , \mathbf{G}_1 , and \mathbf{G}_2 at each node. These three sparse matrices are then extracted from R-INLA, and used in Template Model Builder (TMB, Kristensen et al. 2013) in subsequent steps of the maximum likelihood estimation.

Maximum likelihood estimation proceeds by defining $\Omega^{(k)}$ and $\Psi_t^{(k)}$, i.e., random fields defined at each node in the SPDE approximation. These follow a multivariate normal distribution.

$$\begin{aligned}\Omega &\sim MVN(\alpha\mathbf{1}, \Sigma_\Omega) \\ \Psi &\sim MVN(0, \Sigma_E \otimes \Sigma_\psi)\end{aligned}\tag{B.2}$$

where:

$$\begin{aligned}\Sigma_\Omega &= (\tau_\Omega^2(\kappa_\Omega^4 \mathbf{C} + 2\kappa_\Omega^2 \mathbf{G}_1 + \mathbf{G}_2))^{-1} \\ \Sigma_E &= (\tau_E^2(\kappa_E^4 \mathbf{C} + 2\kappa_E^2 \mathbf{G}_1 + \mathbf{G}_2))^{-1}\end{aligned}\tag{B.3}$$

and Σ_ψ is as defined in Eq. 2c (in the present application, we assume that $\kappa_\Omega = \kappa_E$, although future applications could explore the consequences of relaxing this assumption). $\Omega^{(k)}$ and $\Psi_t^{(k)}$ at knots that correspond to stations with data are then used to calculate the conditional probability of available data.

The computational cost of this SPDE approximation is $O(n^{3/2})$, while the cost of inverting the original Gaussian random field is $O(n^3)$, so this approximation is expected to gain in importance as the number of stations for available data increases. Following an empirical hierarchical modelling strategy (*sensu* Cressie and Wikle 2011), $\Omega^{(k)}$ and $\Psi_t^{(k)}$ are integrated across while calculating the marginal likelihood of κ , τ_Ω , τ_E , α , and any other hyperparameters of interest. We then use the delta-method to back-calculate the value of interpretable parameters, i.e., the distance at which the correlation has fallen to approximately 13% of its maximum (the spatial “range” λ):

$$\lambda = \frac{\sqrt{8\nu}}{\kappa}\tag{B.4}$$

and the marginal variance σ^2 of the random field:

$$\sigma^2 = \frac{\Gamma(\nu)}{\Gamma(\nu + 0.5d)(4\pi)^{d/2} \kappa^{2\nu} \tau^2} \quad (\text{B.5})$$

where Γ is the gamma function, d is the dimension (i.e., 2 in the 2-dimensional spatial model), $\nu=1$ as assumed in the Matérn approximation, and κ and τ are the estimated parameters.

LITERATURE CITED

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